Hierarchical Models in Logistic Regression

Motivation by Example

Suppose that we have a data set where nine different MD's made decisions on 133 patients in total. On average, about 15 patients per MD, although the data were not evenly distributed across MDs. The decisions involved whether to give a patient thrombolysis or not in the emergency room.

There are Canadian guidelines as to whether to thrombolyse such patients or not, so each physician was evaluated as to whether they followed the guidelines for each subject they treated or not. For example, the first physician followed the guidelines in 19 out of 20 patients she/he treated. While each physician has their own rate of "success" (following the guidelines), it may be that overall, these rates may themselves follow a pattern, i.e., have a statistical distribution.

The question is:

How can we best estimate each physician's rate of correct decision making?

If we use each physician's rate separately, then the sample sizes will be too small.

For example, the largest number of patients contributed by any MD was 24, and so we would expect a very wide interval estimate from even the best case scenario.

On the other hand, it is probably quite unreasonable to assume that the true rate of correct decision making is constant across physicians, as each has their own training, background, and so on. Thus, we expect rates to different across physicians.

So, pooling of all data would provide an accurate estimate, but ignores individual variability. Individual estimates take each MD's ability into account, but will lead to very poor accuracy owing to small sample sizes.

Is there a way to address this problem?

One answer is to use a hierarchical/random effects model, which allows us to reasonably accurately estimate both individual and overall rates of guideline adherence, in exchange for an assumption: That the nine MDs are like a sample from a group of MDs that collectively have rates that follow a statistical distribution. This "higher level" distribution is why these models are called "hierarchical models".

The data are from the MSc thesis of Dr. Michael Schull.

Model We will now program this problem using a hierarchical model in WinBUGS.

```
model                    #  Usual model statement in WinBUGS
{
 for (i in 1:nmd) {      #  Loop over number of MDs = 9
 x[i] ~ dbin(p[i],n[i])  #  For each MD, number of
                         #  successes is binomial
                         #  with size = n[i], rate = p[i]

  logit(p[i]) <- z[i]    #  Create a hierarchical
 z[i] ~ dnorm(mu,tau)    #  distribution, by letting
      }                  #  logit of rates be normally
                         #  distributed
                         #  z is a "transformed" variable
                         #  representing the rates p[i]
                         #  on the logit scale

    mu ~  dnorm(0,0.001)      # Prior distribution for mu
   tau ~ gamma(0.001, 0.001) # Prior distribution for tau
   y   ~ dnorm(mu, tau)       #  Predictive distribution for rate on logit scale
 sigma <- 1/sqrt(tau)          #  Prior SD on the logit scale, FROM TAU
    w <- exp(y)/(1+exp(y))    #  Predictive dist transformed back to prob scale
}
```

Data

list(n=c(20, 6, 24, 13, 12, 4, 24, 12, 18),
x=c(19, 5, 22, 12, 11, 4, 23, 12, 16),
nmd=9)

Initial Values

list(mu=0, sigma=1)

Results

node	mean	sd	MC error	2.5%	median	97.5%	start	sample
mu	2.642	0.3716	0.0294	1.973	2.63	3.357	1001	5000
tau	215.9	403.8	21.61	1.189	51.67	1447.0	1001	5000
sigma	0.2245	0.2467	0.01485	0.02631	0.1393	0.9187	1001	5000
p[1]	0.9301	0.02673	0.0018	0.8693	0.9332	0.9718	1001	5000
p[2]	0.9242	0.03447	0.002022	0.8463	0.9294	0.9689	1001	5000
p[3]	0.9277	0.0266	0.001763	0.8665	0.9316	0.9684	1001	5000
p[4]	0.9281	0.02856	0.001852	0.8598	0.9321	0.9712	1001	5000
p[5]	0.9279	0.0286	0.001851	0.8635	0.9319	0.9721	1001	5000
p[6]	0.9293	0.02987	0.001896	0.8636	0.9333	0.9758	1001	5000
p[7]	0.9313	0.02574	0.001803	0.8745	0.9343	0.9739	1001	5000
p[8]	0.932	0.02688	0.001838	0.8748	0.9347	0.9779	1001	5000
p[9]	0.9257	0.0288	0.001872	0.8595	0.93	0.9677	1001	5000
w	0.927	0.03535	0.001954	0.8524	0.9324	0.9725	1001	5000
y	2.638	0.4917	0.02993	1.753	2.624	3.565	1001	5000

Rerunning above, but with data changed so that first row is 10/20. Note how this "outlier" is not pulled back as strongly towards the overall mean.

node	mean	sd	MC error	2.5%	median	97.5%	start	sample
mu	2.347	0.6308	0.02193	1.315	2.286	3.806	1001	4000
p[1]	0.5969	0.1161	0.003224	0.366	0.6	0.8094	1001	4000
p[2]	0.8614	0.0981	0.001751	0.6063	0.8834	0.9825	1001	4000
p[3]	0.9076	0.05137	0.001297	0.7848	0.9152	0.9808	1001	4000
p[4]	0.9075	0.06156	0.001603	0.7596	0.9193	0.9889	1001	4000
p[5]	0.9027	0.06282	0.001518	0.7504	0.9153	0.9875	1001	4000
p[6]	0.9104	0.07954	0.001975	0.6975	0.9322	0.9972	1001	4000
p[7]	0.933	0.04373	0.001305	0.8296	0.9413	0.9918	1001	4000
p[8]	0.9404	0.05143	0.001714	0.8086	0.9537	0.9982	1001	4000
p[9]	0.8888	0.06163	0.001315	0.7372	0.8989	0.9775	1001	4000
sigma	1.205	0.5925	0.02313	0.366	1.106	2.64	1001	4000
tau	1.525	2.806	0.1283	0.1441	0.8176	7.476	1001	4000
w	0.8579	0.1527	0.002682	0.385	0.9027	0.9957	1001	4000
y	2.339	1.452	0.03215	-0.4683	2.228	5.446	1001	4000

Rerunning above, but with data changed so that first row is 100/200. Now very strong outlier, results stays near 50% observed rate.

node	mean	sd	MC error	2.5%	median	97.5%	start	sample
mu	2.399	0.6767	0.01764	1.236	2.345	3.935	1001	4000
p[1]	0.5088	0.03571	5.547E-4	0.4406	0.5092	0.5817	1001	4000
p[2]	0.8565	0.1035	0.002056	0.5942	0.8801	0.9838	1001	4000
p[3]	0.9104	0.05155	9.935E-4	0.7889	0.9198	0.9822	1001	4000
p[4]	0.91	0.06435	0.00122	0.746	0.9248	0.99	1001	4000
p[5]	0.9066	0.06611	0.001272	0.7396	0.9202	0.9895	1001	4000
p[6]	0.9191	0.08212	0.002	0.6984	0.9439	0.9985	1001	4000
p[7]	0.9397	0.04166	8.286E-4	0.8367	0.9485	0.9934	1001	4000
p[8]	0.9499	0.04702	0.001143	0.8217	0.9635	0.9989	1001	4000
p[9]	0.8894	0.06441	0.001169	0.7325	0.902	0.9784	1001	4000
sigma	1.433	0.6118	0.01963	0.6696	1.311	3.024	1001	4000
tau	0.7414	0.5728	0.01634	0.1097	0.5827	2.242	1001	4000
w	0.8457	0.1816	0.003351	0.2951	0.9117	0.9977	1001	4000
y	2.396	1.714	0.03377	-0.8707	2.335	6.091	1001	4000

So, to summarize:

The hierarchical model allowed one to estimate the rate of each physician reasonably accurately, assuming the model is reasonable.

When all rates were similar, sigma is small, and rates are bunched together (a lot of "borrowing of strength"). When one (or in general, any number) rate sticks out from the rest, the degree of "pull" back towards the overall mean depends on how strong the result is (which in turn, depends on how different it is, and on the sample sizes. Note that sigma grows in size as the individual rates are spread further apart, meaning less "pull" back towards the overall mean.

Using this model, can also predict what the overall rate is, as well as the rate for the "next similar physician".

Example 2: Hierarchical Logistic Regression

The above example used a logit function, but there were no covariates. as we saw for linear regression, on can also place hierarchical components on intercepts and beta coefficients.

Consider the following example:

We would like to estimate the rates of osteoporosis in Canada, but believe that they differ not only by age and sex (higher rates in older females compared to younger males), but also by province.

We have a data set of 10,000 female subjects, 1000 from each province. For each subject, we have their age, province, and whether they have been diagnosed with osteoporosis.

The data are in a file called osteo.txt, and we will create a hierarchical model that accounts for possible differences across provinces, and provides probabilities that one province has a higher rate compared to any other province.

The data structure is as follows:

The first 1000 data points are from Newfoundland, the next 1000 from Nova Scotia, the next 1000 from New Brunswick, and then PEI, Quebec, Ontario, Manitoba, Saskatchewan, Alberta, and BC, in that order.

```
model                           #  Usual model statement in WinBUGS
{
 for (j in 1:10)                #  Loop over 10 provinces
{
                                #  Common index trick
 for (i in index[j]:index2[j]) # Index for jth province
     {
     logit(p[i]) <- alpha[j] + beta*age[i] #  Logit for individual probability
     osteo[i] ~ dbern(p[i])                #  Likelihood function for ith individual
     }                                    #
     alpha[j] ~ dnorm(mu, tau)          #  Hierarchical component:  provincial rates
   }                                    #  "tied together" through normal distribution
 mu ~  dnorm(0,0.001)                   #  Prior on hierarchical mean
 tau <- 1/(sigma*sigma)                 #  Needed for WinBUGS
sigma ~ dunif(0,20)                     #  Prior for hierarchical Sd
 beta  ~  dnorm(0, 0.001)               #  Prior for beta
pred.NFLD.50 <- exp(alpha[1] + beta*50)/(1+exp(alpha[1] + beta*50))
pred.QUEBEC.50 <- exp(alpha[5] + beta*50)/(1+exp(alpha[1] + beta*50))
pred.BC.50 <- exp(alpha[10] + beta*50)/(1+exp(alpha[10] + beta*50))
diff.quebec.bc <-  pred.QUEBEC.50 - pred.BC.50  # Quebec - BC diff
p.quebec.bc <-  step(pred.QUEBEC.50 - pred.BC.50)  # p(Quebec > BC)
}

# Inits

list(alpha=c(0,0,0,0,0,0,0,0,0,0), beta=0.5, mu=0, sigma = 1)
```

```
#  Data

list(index = c(1, 1001, 2001, 3001, 4001, 5001, 6001, 7001,
8001, 9001),
index2 =c(1000, 2000, 3000, 4000, 5000, 6000, 7000,
8000, 9000, 10000),
age=c(39, 25, 36, 68, 67, 50, 39, 68, 31, 30, 54, 30, 30,
71, 30, 28, 49, 54, 29, 61, 35, 52, 29, 74, 70, 50, 41, 74, 74,
38, 29, 41, 27, 28, 57, 32, 32, 64, 45, 60, 56, 66, 65, 40, 64,
58, 51, 30, 48, 63, 50, 52, 67, 47, 44, 37, 41, 32, 35, 27, 58,
.....................etc.................
55, 40, 27, 33, 37, 39, 28, 75, 37, 52, 57, 67, 25, 33, 34, 55,
39, 46, 28), osteo=c(0, 0, 0, 0, 0, 1, 0, 1, 0, 0, 1, 0, 0, 0, 1, 0, 0,
0, 0, 0, 1, 1, 1, 0, 1, 0, 1, 1, 0, 0, 0, 1, 1, 0, 0, 0, 0, 0,
0, 0, 0, 1, 1, 0, 1, 0, 0, 0, 1, 0, 0, 0, 1, 0, 1, 1, 0, 0, 0,
0, 0, 0, 1, 0, 1, 1, 0, 1, 0, 0, 0, 1, 0, 1, 0, 1, 0, 0, 0, 0,
0, 1, 1, 1, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 1, 1, 1, 0, 1,
.....................etc.................
0, 1, 0, 0, 0, 0, 0, 1, 1, 0, 0, 1, 0, 0, 0, 0, 1, 0, 1, 1, 0,
1, 0, 1, 0, 0, 0, 0, 1, 1, 1, 1, 0, 0, 0, 1, 0, 0, 0, 0, 1, 1,
0, 0, 1, 1, 1, 1, 1, 1))
```

The results are as follows:

node	mean	sd	MC error	2.5%	median	97.5%	start	sample
alpha[1]	-0.6307	0.09387	0.003661	-0.8123	-0.6306	-0.4466	1001	10000
alpha[2]	-0.5916	0.09053	0.00362	-0.7634	-0.5923	-0.414	1001	10000
alpha[3]	-0.437	0.08762	0.003624	-0.6053	-0.4376	-0.2607	1001	10000
alpha[4]	-0.4934	0.08708	0.003542	-0.6622	-0.4931	-0.3201	1001	10000
alpha[5]	-0.4047	0.0882	0.003568	-0.5745	-0.4054	-0.2329	1001	10000
alpha[6]	-0.3533	0.09058	0.003677	-0.5269	-0.3532	-0.1722	1001	10000
alpha[7]	-0.3739	0.09001	0.003639	-0.5478	-0.3743	-0.1951	1001	10000
alpha[8]	-0.4318	0.08846	0.003605	-0.6019	-0.433	-0.2543	1001	10000
alpha[9]	-0.4584	0.0879	0.003628	-0.627	-0.4588	-0.2862	1001	10000
alpha[10]	-0.5	0.087912	0.003544	-0.6734	-0.5013	-0.3295	1001	10000
beta	0.002226	0.001352	6.719E-5	-4.721E-4	0.002239	0.004797	1001	10000
diff.quebec.bc	0.06421	0.03328	4.504E-4	0.003522	0.06245	0.1322	1001	10000
mu	-0.4677	0.08147	0.003604	-0.6229	-0.4686	-0.3052	1001	10000
p[1]	0.3674	0.01559	3.223E-4	0.3369	0.3675	0.3984	1001	10000
p[2]	0.3602	0.01694	4.869E-4	0.3275	0.3602	0.394	1001	10000
p[3]	0.3659	0.01579	3.533E-4	0.3348	0.3659	0.3973	1001	10000
p[4]	0.3825	0.01654	3.561E-4	0.3499	0.3826	0.415	1001	10000

p[9998]	0.3982	0.014	2.502E-4	0.3705	0.3983	0.4261	1001	10000
p[9999]	0.4019	0.01366	1.677E-4	0.3749	0.402	0.4289	1001	10000
p[10000]	0.3924	0.0152	4.08E-4	0.3623	0.3923	0.4223	1001	10000
p.quebec.bc	0.9823	0.1319	0.001347	1.0	1.0	1.0	1001	10000
pred.BC.50	0.4041	0.01363	1.404E-4	0.377	0.4041	0.4311	1001	10000
pred.NFLD.50	0.3731	0.01532	2.523E-4	0.3428	0.3731	0.4033	1001	10000
pred.QUEBEC.50	0.4683	0.03007	3.949E-4	0.4144	0.4666	0.5307	1001	10000
sigma	0.1212	0.04609	8.6E-4	0.05246	0.114	0.2299	1001	10000
tau	172.6	1786.0	69.06	18.95	76.94	364.7	1001	10000

One could have predicted rate for next person from each province of a given age, and so on. Note that sigma was relatively large (on the logit scale) indicating substantial differences in rates between provinces. Note that Quebec is highly likely to have a larger rate than BC, with probability = 0.98. Also, note the very direct statement arising from this variable, in comparison with p-values.

Final Comments

- One can have multiple level hierarchical models, accounting for hierarchical data structure. For a detailed example see the paper:

  Brophy J, Joseph L, Theroux P, on behalf of the Quebec Acute Coronary Care Working Group. Medical decision making about the choice of thrombolytic agent for Acute Myocardial Infarction. Medical Decision Making 1999;19(4):411-418.

  This paper used a four level model. It was programmed in WinBUGS, and formed part of the PhD thesis of Dr. James Brophy.

- Frequentist random effects models can be programmed in R, perhaps less intuitive than those in WinBUGS, beyond the scope of this course.

- As discussed for linear regression models, one can almost always argue for a hierarchial component to any regression model in epidemiology: Subjects are treated by different physicians, in different hospitals, at different times, in different provinces, and so on.

  Overall, hierarchical modelling is a very important topic, we have just scratched the surface here, but be on the lookout for opportunities to better your models through addition of hierarchical levels.